1	Machine Learning for the ECG Diagnosis and Risk Stratification of
2	Occlusion Myocardial Infarction at First Medical Contact
3	Results from ECG-SMART Observational Trial
4	Ву
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30 ABSTRACT

- 31 Patients with occlusion myocardial infarction (OMI) and no ST-elevation on presenting ECG are
- 32 increasing in numbers. These patients have a poor prognosis and would benefit from immediate
- 33 reperfusion therapy, but we currently have no accurate tools to identify them during initial triage. Herein,
- 34 we report the first observational cohort study to develop machine learning models for the ECG
- diagnosis of OMI. Using 7,313 consecutive patients from multiple clinical sites, we derived and
- 36 externally validated an intelligent model that outperformed practicing clinicians and other widely used
- 37 commercial interpretation systems, significantly boosting both precision and sensitivity. <u>Our derived</u>
- 38 OMI risk score provided enhanced rule-in and rule-out accuracy relevant to routine care, and when
- 39 combined with the clinical judgment of trained emergency personnel, this score helped correctly
- 40 reclassify one in three patients with chest pain. ECG features driving our models were validated by
- 41 clinical experts, providing plausible mechanistic links to myocardial injury.

42	The ECG diagnosis of acute coronary syndrome (ACS) in patients with acute chest pain is a
43	longstanding challenge in clinical practice. ¹⁻⁴ Guidelines primarily focus on ST-segment elevation (STE)
44	for discerning patients with ST-elevation myocardial infarction (STEMI) vs. other forms of ACS. ⁵⁻⁸ A
45	biomarker-driven approach is recommended in the absence of STE on the presenting ECG. This
46	diagnostic paradigm has two important limitations. First, around 24%-35% of patients with non-STEMI
47	have total coronary occlusion, referred to as occlusion myocardial infarction (OMI), and require
48	emergent catheterization.9-13 This vulnerable group, in contrast to acute myocardial infarctionACS with
49	an open artery (non-OMI) (Extended Data Fig. 1), suffers from unnecessary diagnostic and treatment
50	delays that are associated with higher mortality. ¹⁴⁻¹⁷ This excess risk can be mitigated with enhanced
51	diagnostic criteria. Although important ECG signatures of OMI are frequently described in the
52	literature, ¹⁸⁻²¹ they are subtle, involve the entire QRST complex, and are spatial in nature (i.e., changes
53	diluted across multiple leads). ²²⁻²⁴ Visual inspection of ECG images by clinical experts, thus, is
F 4	suboptimal and leads to a high degree of variability in ECG interpretation. ²⁵⁻²⁷
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67 In our prior work, we designed prototype algorithms for AI-enabled ECG analysis and 68 demonstrated the clinical feasibility of screening for ACS in the prehospital setting.^{34,35} Herein, we 69 describe the first multisite, prospective, observational cohort study to evaluate the diagnostic accuracy 70 of machine learning for the ECG diagnosis and risk stratification of OMI at first medical contact in an 71 observer-independent approach (Extended Data Fig. 2). Our intelligent models were derived and 72 externally validated on 7,313 patients with chest pain from multiple clinical sites in the United States. 73 The results demonstrate the superiority of machine learning in detecting subtle ischemic ECG changes indicative of OMI, outperforming practicing clinicians and other widely used commercial ECG 74 interpretation software. Our derived OMI risk score provides superior-enhanced rule-in and rule-out 75 76 accuracy when compared to the HEART score, helping correctly reclassify one in three patients with 77 chest pain. We identified the most important ECG features driving our model's classifications and 78 identified plausible mechanistic links to myocardial injury.

79

80 RESULTS

81 Sample Characteristics

After excluding patients with cardiac arrest, ventricular tachyarrhythmias, confirmed prehospital 82 83 STEMI, and duplicate ECGs, our derivation cohort included 4,026 consecutive patients with chest pain (age 59±16 years, 47% females, 5.2% OMI). The two external validation cohorts together included 84 3,287 patients (age 60±15 years, 45% females, 6.4% OMI) (Fig. 1 and Table 1). Most patients in the 85 derivation and validation cohorts were in normal sinus rhythm (>80%) and around 10% were in atrial 86 fibrillation. Around 3% of patients had left bundle branch block (LBBB) and ~10% had ECG-evidence of 87 88 left ventricular hypertrophy (LVH). The derivation and validation cohorts were comparable in terms of 89 age, sex, baseline clinical characteristics, and 30-day cardiovascular mortality. The validation cohort, however, had more Black and Hispanic minorities and a slightly higher rate of ACS and OMI. The 90

91	presence of OMI, defined as a culprit coronary artery with a TIMI flow grade of 0-1, was adjudicated
92	from charts by independent reviewers blinded to all ECG analyses. A TIMI flow grade of 2 with
93	significant coronary narrowing (>70%) and peak 4 th generation (not high sensitivity) troponin of 5-10
94	ng/mL was also indicative of OMI.
95	Algorithm Derivation and Testing
96	The positive class for model training was the presence of OMI, defined as a culprit coronary
97	artery with a TIMI flow grade of 0-1, as adjudicated from charts by independent reviewers blinded to all
98	ECG analyses. A TIMI flow grade of 2 with significant coronary narrowing (>70%) and peak 4th
99	generation (not high sensitivity) troponin of 5-10 ng/mL was also indicative of OMI. The negative class
100	for model training was the absence of OMI, which included all other non-ACS etiologies or those with
101	non-coronary occlusive ACS subtypes.
102	Input data for model training was based on prehospital 12-lead ECGs obtained at first medical
103	contact. We selected 73 morphological ECG features out of 554 temporal-spatial metrics using a hybrid
104	data-driven and domain expertise approach. ¹⁸ Using these features, ten classifiers were trained to learn
105	ischemic patterns between ACS and non-ACS groups and to estimate the probability of OMI:
106	regularized logistic regression, linear discriminant analysis, support vector machine, Gaussian Naïve
107	Bayes, random forest, gradient boosting machine, extreme gradient boosting, stochastic gradient
108	descent logistic regression, k-nearest neighbors, and artificial neural networks. We chose these
109	classifiers because they learn different mathematical representations in the data, in order to
110	maximiz <u>eing</u> the chance of finding the best <u>fitting</u> modelingapproach approach for <u>learning the</u>
111	mathematical representation relating complex ECG data to underlying physiology.
112	The random forest model achieved the best bias-variance tradeoff for training and internal
113	testing. We compared the random forest against the ECG interpretation of practicing clinicians and
114	against the performance of a commercial ECG interpretation system that is FDA-cleared for "Acute MI"
115	diagnosis. On the hold-out test set, the random forest model (AUROC 0.91 [95% CI 0.87-0.96])

outperformed both practicing clinicians (AUROC 0.79 [95% CI 0.73-0.76], p<0.001) and the commercial

117 ECG system (AUROC 0.78 [95% CI 0.70-0.85], p<0.001) (Fig. 2A).

Next, we used probability density plots for OMI(+) and OMI(-) classes to denote the optimal 118 separation margins for risk prediction. As recommended by guidelines,⁶ we defined a risk score to 119 120 identify patients at low risk (OMI score <5), intermediate risk (OMI score 5-20), and high risk (OMI score >20), with these cutoffs yielding excellent separation between classes (Log-rank chi-square 121 133.04, df=2, p<0.001) (Fig. 2B, left panel). Our OMI score classified 74.4% of patients as low-risk and 122 123 4.6% as high-risk. Using the low-risk group in a rule-out strategy yielded a sensitivity of 0.91 and a negative predictive value (NPV) of 0.993, with an overall missed event rate of 0.5%. Using high-risk 124 125 class for a rule-in strategy yielded a specificity of 0.976 and a positive predictive value (PPV) of 0.514, with an overall false discovery rate of 2%. Finally, we compared this OMI score to the HEART score, 126 which uses patient history, ECG data, age, risk factors, and troponin values (Fig. 2B, right panel). Our 127 128 OMI score, which is based on ECG data alone, classified 66% more patients as low risk than the 129 HEART score with a comparable false negative rate <1%, and classified fewer patients as high-risk and with much higher precision (51% vs. 33%). The OMI score also triaged 50% fewer patients as 130 intermediate risk and still got better discrimination for OMI detection (11.2% vs. 5.6%). 131 132 Model Explainability 133 We used Tree SHAP algorithms to generate an importance ranking that explains the output of the 134 random forest model based on SHAP values estimated for the top 25 features (Fig. 3A). The features with the greatest impact on classification output included slight ST-depression in leads V1, V2, I, and 135 aVL; slight ST-elevation in leads III and V4-V6; loss of concave pattern in anterior leads; T wave 136 137 enlargement in II and aVF and T flattening or inversion in I and aVL; prolonged T_{peak}-T_{end} interval; T axis deviation; increased repolarization dispersion; and distorted directions of activation and recovery 138

139 patterns. Most of these ECG patterns can be mechanistically linked to cardiac ischemia, suggesting

140 their clinical value as plausible features for OMI detection.

142	To better visualize these global ECG patterns detected by our model, we created pooled
143	population median beats for the OMI(+) class (n=414 ECGs), and superimposed these median beats
144	on the pooled population median beats of patients with normal sinus rhythm and OMI(-) status (n=9,072
145	ECGs) (Fig. 3B). Findings from this figure agree with the patterns derived from the SHAP values
146	described above. Specifically, this figure illustrates that OMI is associated with ST-depression and T $$
147	flattening in V1-V2, I, and aVL; slight ST-elevation in the anterior leads with loss in concave pattern;
148	peaked T wave in inferior leads; T_{peak} - T_{end} prolongation (seen in many leads); global repolarization
149	dispersion (seen as peaked T in some leads and flattening in others); T axis deviation (away from the
150	left ventricle), and distorted activation and recovery patterns (seen in the horizontal plane as loss of R
151	wave progression in precordial leads with increased T wave discordance). Due to prevalent multivessel
152	disease in this cohort, these OMI patterns remained relatively consistent regardless of culprit location.
153	Nevertheless, to examine local explainability of feature importance, we used force plots on
154	individual cases to identify the features that met the contribution threshold of the random forest model
155	on a given ECG. These force plots were also examined by study investigators to further corroborate on

the clinical validity of model predictions. Extended Data Fig. 3 shows a selected example of a 12-lead

157 <u>ECG with its corresponding force plot for the local features contribution.</u>

158 External Validation

141

We tested the final lock-out model on 3,287 patients from two independent external clinical sites. Machine learning engineers were blinded to outcome data from other sites, and the pre-populated model predictions were independently evaluated by the clinical investigators. Our model generalized well and maintained high classification performance (AUROC 0.873 [95% CI 0.85-0.90]), outperforming the classification performance of the commercial ECG system (AUROC 0.75 [95% CI 0.71-0.79], p<0.001) and practicing clinicians (AUROC 0.80 [95% CI 0.77-0.83], p<0.001) (**Fig. 4A**). Our OMI risk score was a strong predictor of OMI, independent from, age, sex, and other coronary risk factors (OR Formatted: Font: Bold

166 10.6 [95% CI 6.78-16.64] for high-risk class and OR 2.85 [95% CI 1.91-4.28] for intermediate-risk class) 167 (Fig. 4B). This risk score triaged 69% of patients in the low-risk group at a false-negative rate of 1.3% 168 and identified 5.1% of patients as high-risk at acceptable true positive rate >50%. The overall 169 sensitivity, specificity, PPV, and NPV for the OMI rule-in and rule-out strategy were 0.86 (95% CI 0.81-170 0.91), 0.98 (95% CI 0.97-0.99), 0.54 (95% CI 0.46-0.62), and 0.99 (95% CI 0.98-0.99), respectively. 171 This diagnostic accuracy remained relatively similar across subgroups based on age, sex, race, 172 comorbidities, and baseline ECG findings, indicating the lack of aggregation bias (Fig. 4C). In comparison, the sensitivity, specificity, PPV, and NPV for ECG overread by practicing clinicians were 173 0.58, 0.93, 0.36, and 0.97, and for the commercial ECG system 0.79, 0.80, 0.22, and 0.98, respectively. 174 Next, we used decision analysis to evaluate the incremental gain of our derived risk score in re-175 classifying patients at first medical contact (Fig. 5). To simulate initial assessment by emergency 176 personnel - we used was based on the modified HEAR score (History, ECG, Age, and Risk factors) to 177 178 triage patients into low, intermediate, and high-risk groups.³⁶ At baseline, emergency personnel triaged 179 48% of patients as low risk with a NPV of 99.0% and triaged 3% of patients as high risk with a PPV of 54.1%. Nearly 50% of patients remained in an indeterminate observation zone. Applying our OMI risk 180 181 score would help triage 45% more patients as low risk while keeping the NPV at 98.8% and would help detect 85% more cases with OMI while keeping PPV at 50.0%. The OMI score would also help reduce 182 the number of patients in the indeterminate observation zone by more than half. These numbers 183 translate into a net reclassification improvement (NRI) index of 41% (95% CI 33%-50%). To validate 184 this incremental clinical utility, we manually reviewed ECGs reclassified correctly as OMI(+) (Extended 185 data Fig. 34). Many of these ECGs showed subtle or nonspecific changes that were nondiagnostic as 186 187 per guidelines,⁵ suggesting potential value in boosting provider's confidence when interpreting "fuzzy" ECGs. 188

Finally, we investigated the potential sources of false negatives in the validation data. Among those with missed OMI events (n=28, 0.9%), many patients had high-frequency noise and baseline

191	wander on their initial ECG (n=13/28, 46%) or had low voltage ECG (n=14/28, 50%), and most patients
192	(n=24/28, 86%) had benign ECGs without any diagnostic ST-T changes (Extended Data Fig. 4 <u>5</u>).
193	Moreover, we found no significant differences between false negatives and true positives in terms of
194	demographics or clinical characteristics, with the exception that most false negatives had a history of a
195	prior myocardial infarction (93% vs. 27%). The latter finding was intriguing given that our OMI model
196	was slightly less specific in patients with known coronary artery disease (Fig. 4C). This finding aligns
197	with recent evidence showing diminished NPV in patients with chest pain and known CAD.37

198 Screening for Any ACS Event

199 We further built a model to screen for any potential ACS event at first medical contact. Using the same set of ECG features, we trained and optimized a random forest classifier that denotes the 200 201 likelihood of any ACS event. The model performed well during training (AUROC 0.88 [95% CI 0.87-0.90]) and generalized well during internal testing (AUROC 0.80 [95% CI 0.76-0.84]), outperforming 202 both the commercial ECG interpretation system (AUROC 0.62 [95% CI 0.55-0.68], p<0.001) and 203 practicing clinicians (AUROC 0.66 [95% CI 0.59-0.72], p<0.001) (Extended Data Fig. 56). On external 204 validation, the model continued to generalize well (AUROC 0.79 [95% CI 0.76-0.81]), outperforming the 205 commercial system (AUROC 0.68 [95% CI 0.65-0.71], p < 0.001) and practicing clinicians (AUROC 206 207 0.72 [95% CI 0.69-0.74], p < 0.001). Our derived risk score provided a suboptimal rule-out classification for any ACS event (sensitivity 68.2% and NPV 92.5%) but provided superior rule-in accuracy 208 (specificity 98.9% and PPV 82.5%). 209

210 **DISCUSSION**

In this study, we developed and validated a machine learning algorithm for the ECG diagnosis of OMI in consecutive patients with chest pain recruited from multiple clinical sites in the United States. This model outperformed practicing clinicians and other commercial interpretation systems. The derived risk score provided superior rule-in and rule-out accuracy for OMI, boosting the sensitivity by 7 to 28 percentage points and the precision by 18 to 32 percentage points compared to reference standards.

216	When combined with the judgment of experienced emergency personnel, our derived OMI risk score
217	helped correctly reclassify one in three patients with chest pain. To our knowledge, this is the first study
218	using machine learning methods and novel ECG features to optimize OMI detection in patients with
219	acute chest pain and negative STEMI pattern on their baseline ECG at first medical contact.
220	Mapping myocardial ischemia, a problem of regional metabolic derangement, to coronary
221	occlusion, a problem of diminished blood flow due to an atherosclerotic plaque rupture, is a complex
222	process. ¹ Essentially, ischemia disproportionately distorts action potentials in different myocardial
223	segments, resulting in tissue-scale currents, often called 'injury' currents. Prior studies have mapped
224	significant ST-elevation to transmural injury currents associated with total coronary occlusion. This has
225	historically driven the current paradigm dichotomy of STEMI vs. 'others' (any ACS other than STEMI) in
226	determining who might benefit from emergent reperfusion therapy. However, nearly 65% of patients
227	with ACS present with no ST-elevation on their baseline ECG, ^{35,38} and among the latter group,
228	24%-35% have total coronary occlusion requiring emergent catheterization.9-13 Thus, determining who
229	would benefit from reperfusion therapy remains an adjudicated diagnosis.
230	Conceptually, injury currents produced by ischemic cardiac cells are summative in nature,
231	explaining how ST amplitude changes can get attenuated on the surface ECG (Extended Data Fig.
232	67). These injury currents, however, distort the propagation of both excitation and recovery pathways,
233	altering the configuration of the QRS complex and the STT waveform altogether. ³⁹ Thus, a more
234	comprehensive approach for the ECG detection of ischemia should focus on (1) evaluating temporal
235	characteristics over entire waveform segments rather than the voltage at a given time point (e.g., J+80),
236	and (2) evaluating lead-to-lead spatial characteristics in waveform morphology rather than absolute
237	changes in isolated ECG leads. ¹
238	This study has identified several ECG patterns indicative of acute coronary occlusion beyond
239	the criteria recommended by clinical guidelines. ⁵ Intriguingly, these ECG patterns overlap with those

240 described in the literature. A consensus report in 2012 identified few ECG patterns that should be

241 treated as STEMI equivalent during acute pain episodes: ST-depression in V1 to V3; small inverted T 242 waves in V1 to V3; deep negative T waves in precordial leads; widespread ST-depression, and 243 prominent positive T waves.²⁰ Similar ECG patterns were also described more recently: ST-depression 244 in V1 to V4 (versus V5-V6); reciprocal ST-depression with maximal ST-depression vector towards the 245 apex (leads II and V5, with reciprocal STE in aVR); subtle ST-elevation; acute pathologic Q waves; 246 hyperacute T waves; and loss of terminal S wave.²¹ Many of these expert-driven patterns rely on 247 assessing the proportion of repolarization amplitudes or area under the QRS amplitude. They also rely heavily on the visual assessment of waveform morphology and can introduce a high degree of 248 subjectivity and variability among ECG interpreters. We demonstrated that the machine learning 249 250 models described herein not only outperform practicing clinicians in identifying OMI, but also provided 251 an objective, observer-independent approach to quantify subtle ECG patterns associated with OMI. Many of the data-driven features identified by our machine learning model are subtle and cannot 252 be easily appreciated by clinical experts. T feature indices were among these most important features, 253 254 including T_{peak}-T_{end} interval prolongation, T wave flattening, and T wave characteristics at the inflection 255 point preceding T_{peak} (Fig. 3A). Mechanistically, ischemic injury currents interfere with signal 256 propagation leading to longer activation time.⁴⁰ These late activation potentials lead to a loss of terminal S wave and longer recovery time, both manifesting as T wave flattening, shifted T peak, and loss of 257 258 concavity at the initial T wave (Fig. 3B). These STEMI-equivalent patterns were previously described in the literature as small or negative T waves with widespread ST-depression or subtle ST- elevation.^{20,21} 259 Another important subtle feature identified by our model was increased ventricular repolarization 260 dispersion, measured using the ratio between the principal components of the STT waveforms (i.e., 261 262 PCA metrics), the direction of the T axis, and the angle between activation and recovery pathways 263 (e.g., total-cosine-R-to-T). Injury currents disproportionately affect the duration and velocity of repolarization across different myocardial segments,⁴¹ resulting in lead-to-lead variability in the 264 morphology of the STT waveform.^{22-24,39,42} These high-risk ECG patterns were previously described as 265 266 a mixture of deep negative T waves and prominent / hyperacute T waves or reciprocal T wave

changes.^{20,21} Despite their subtle nature, our machine learning model provided a more comprehensive, 267 268 quantitative approach to evaluating this inter-lead variability in repolarization morphology. Machine learning is well-suited to address many challenges in 12-lead ECG interpretation. 269 270 Myocardial ischemia distorts the duration and amplitude of the Q wave, R peak, R`, QRS complex, ST 271 segment, and T wave, as well as the morphology and configuration of these waveforms (e.g., 272 upsloping, down-sloping, concavity, symmetry, notching, etc.). These distortions are lead-specific yet 273 come with dynamic inter-lead correlations. Thus, ECG interpretation involves many complex aspects and parameters, making it a highly dimensional, decision space problem.¹ Few experienced clinicians 274 275 excel in such pattern recognition,²¹ which explains why so many OMI cases are not reperfused in a 276 timely way; this is also why simple, rule-based commercial systems that use simple regression models are suboptimal for OMI detection. Machine learning algorithms can provide powerful tools to solve such 277 highly dimensional, non-linear mathematical representations found in 12-lead ECG data. 278 Although the literature on machine learning for the ECG diagnosis of coronary disease is 279 ubiquitous, it comes with many serious limitations. First, many studies focused on detecting the known 280 STEMI group or other subtle ACS phenotypes^{34,35,43,44} rather than the critical group without ST-281 elevation, which is not classified as STEMI and is therefore excluded from STEMI databases. Second, 282 most prior work used open-source ECG datasets like PTB and PTB-XL,⁴⁵ which are highly selected 283 284 datasets that focus on ECG-adjudicated diagnoses. Our unique cohorts included unselected, 285 consecutive patients with clinical profiles and disease prevalence like that seen in real-world settings. 286 Third, many studies used a full range of input features based on both ECG data and clinical data elements (e.g., patient history, physical exam abnormalities, laboratory values, diagnostic tests),46-49 287 288 which limits the applicability to real-world settings. Fourth, to our knowledge, most studies used a single derivation cohort for training and testing,⁵⁰ without the use of an independent validation cohort. Finally, 289 prior studies paid little attention to model explainability,⁵¹ shedding little light on novel markers and 290

291	pathways of ischemia than what is already known. Without explanation aids of clinical meaningfulness,
292	machine learning models for ECG interpretation would have limited clinical utility.52
293	This study has important clinical implications. Our machine learning model can be integrated
294	into systems of care for real-time deployment where risk score assignments can be made readily
295	available to clinicians right at time of ECG acquisition. Such enhanced decision support can help
296	emergency personnel identify 85% more patients with critical coronary occlusion despite the absence of
297	a STEMI pattern on the presenting ECG and without any loss in precision. Our models can also help
298	inform care in more than 50% of patients in whom the initial assessment is indeterminate, placing 45%
299	more patients in the low-risk group for OMI without any loss in NPV. This incremental gain in rule-in and
300	rule-out accuracy can help re-allocate critical emergency resources to those in utmost need while
301	optimizing the clinical workflow. This can impact numerous decisions at first medical contact, including
302	targeted prehospital interventions, catheterization lab activation, administration of anti-ischemic
303	therapies, hospital destination decisions, the need for medical consults, referrals for expedited
304	diagnostic testing (e.g., echocardiogram, imaging scans), and early discharge decisions. Furthermore,
305	until now, clinicians never had sensitive nor highly specific tools that would allow the ultra-early
306	identification of OMI in the absence of a STEMI pattern. Such enhanced diagnostics can allow the
307	design and implementation of prospective interventional trials to assess the therapeutic effectiveness of
308	targeted interventions in this vulnerable group (e.g., early upstream P2Y ₁₂ inhibitor administration, ⁵³
309	emergent vs. delayed reperfusion therapy, ⁵⁴ glucose-insulin-potassium infusion, ⁵⁵ etc.).
310	Several limitations merit consideration. First, the engineered features we used for building our
311	models are based on a manufacturer-specific software. There are known discrepancies between
312	manufacturers in ECG preprocessing and metrics computation, which means that our models would
313	need retraining and validation when using different software for ECG signal processing. Alternatively,
314	deep neural networks can be used to directly analyze raw ECG signal without explicit feature
315	engineering. However, these techniques require much larger sample size for model derivation (e.g.,
1	

316	>10,000) and might not yield a meaningful improvement over feature engineering-based machine
317	learning approaches for traditional 12-lead ECG based diagnosis.56_Second, we found slight differences
318	between the derivation and validation cohorts, specifically in terms of disease prevalence and practicing
319	clinicians' accuracy in ECG interpretation. These cohorts came from two different regions in the U.S.,
320	and EMS systems follow state-specific protocols. It is possible that discrepancies in EMS protocols and
321	in-hospital practices resulted in slight differences in the types and proportions of patients that receive
322	prehospital 12-lead ECGs, as well as in their outcome adjudications. Yet, it is reassuring that our
323	models continued to generalize well between the study sites. Third, it is worth noting that our model for
324	screening for "any ACS event" only boosted the performance of the rule-in arm of the derived risk
325	score. This means that a low-risk determination by our model suggests that a given patient would
326	unlikely have OMI, but they might still have a less subtle phenotype of NSTE-ACS that does not require
327	reperfusion therapy. It is likely that serial ECG testing might improve the detection of this group missed
328	events where a patients might switch to a higher risk category in the following hours, ³⁴ but this remains
329	to be confirmed. Coronary occlusion is a dynamic process that evolves over time, so an initial low risk
330	class by our models should not lead to a lower level of active monitoring. Finally, although this study
331	used prospective patients, all analyses were completed asynchronously with patient care. Prospective
332	validation where OMI probabilities and decision support is provided in real time is warranted.
333	In conclusion, we developed and externally validated machine learning models for the ECG
334	diagnosis of OMI in 7,313 patients with chest pain from multiple sites in the United States. The results
335	demonstrated the superiority of machine learning in detecting subtle ischemic ECG changes indicative
336	of OMI in an observer-independent approach. These models outperformed practicing clinicians and
337	commercial ECG interpretation software, significantly boosting both precision and recall. Our derived
338	OMI risk score provided superior enhanced rule-in and rule-out accuracy when compared to HEAR
339	score, and when combined with the clinical judgment of trained emergency personnel, this score helped
340	correctly reclassify one in three patients with chest pain. The ECG features driving our models were

- 341 evaluated, providing plausible mechanistic links to myocardial injury. Future work should focus on the
- 342 prospective validation where OMI probabilities and decision support is provided in real time.

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348 AUTHOR CONTRIBUTION

- 349 SSA, CMG, JZH, SS, ES, and CWC conceived the study, secured funding, and supervised the research. YB and
- 350 SWS advised on the scientific direction of the study. SSA, JZH, ZF, MOA, KKP, and SH supervised dataset
- 351 creation and annotation. SSA, CMG, JZH, ZF, MOA, and SH supervised clinical outcomes adjudication. REG,
- 352 SSA, MA, ZB, PVD, and NR supervised ECG signal processing and feature extraction. SSA, ZB, NR, SMS, and
- ES performed feature engineering, machine learning modeling, statistical analysis, and results interpretation. SSA
- drafted the manuscript. All authors critically revised the manuscript for important intellectual content. All authors
- provided their final approval of the version to be published. All authors are accountable for the work.

356 DECLARATION OF INTERESTS

357 US Patent # 10820822, Owner: University of Pittsburgh, Inventors: SSA, ES, and CWC.

358 DATA AVAILABILITY

- 359 The ECG SMART trial makes use of extracted ECG features to train and evaluate a random forest classifier to
- denote the probability of OMI. The Python codes to evaluate these models along with the derivation and external
- 361 validation datasets are available through GitHub (https://github.com/zeineb-bouzid/sharing-github-nature-
- 362 medicine.git). Researchers wishing the source binary files to design alternative models should contact the
- 363 corresponding author to arrange for proper approvals and institutional data use agreements.

364 CODE AVAILABILITY

- 365 The Python codes to evaluate these models along with the derivation and external validation datasets are
- 366 available through GitHub (<u>https://github.com/zeineb-bouzid/sharing-github-nature-medicine.git</u>).
- 367

368 ONLINE METHODS

369 Ethics Statement

370 The derivation cohort included prehospital data from the City of Pittsburgh Bureau of 371 Emergency Medical Services (EMS) and in-hospital data from three tertiary care hospitals from the University of Pittsburgh Medical Center (UPMC) healthcare system: UPMC Presbyterian Hospital, 372 373 UPMC Shadyside Hospital, and UPMC Mercy Hospital (Pittsburgh, Pennsylvania, USA). All 374 consecutive eligible patients were recruited under a waiver of informed consent. This observational trial 375 was approved by the institutional review board of the University of Pittsburgh and was registered in 376 www.ClinicalTrials.gov (identifier # NCT04237688). The analyses described in this paper were prespecified by the trial protocol that was funded by the National Institute of Health. The first external 377 validation cohort included data from Orange County EMS (Chapel Hill, North Carolina, USA). This study 378 379 actively consented eligible patients and was approved by the institutional review board of the University 380 of North Carolina at Chapel Hill. The second external validation cohort included data from Mecklenburg 381 County EMS and Atrium Health (Charlotte, North Carolina, USA). Data were collected through a 382 healthcare registry and all consecutive eligible patients were enrolled under a waiver of informed 383 consent. This study was also approved by the institutional review board of the University of North 384 Carolina at Chapel Hill. These two external cohorts were very comparable and were, therefore, combined into one cohort. 385

386 Study Design & Data Collection

This was a prospective, observational cohort study. The methods for each study cohort were described in detail elsewhere.^{57,58} All study cohorts enrolled adult patients with an emergency call for non-traumatic chest pain or anginal equivalent symptoms (arm, shoulder, jaw pain, shortness of breath, diaphoresis, syncope). Eligible patients were transported by an ambulance and had at least one recorded prehospital 12-lead ECG. There were no selective exclusion criteria based on sex, race, comorbidities, or acuity of illness. For this prespecified analysis, we only included non-duplicate ECGs
from unique patient encounters, and we removed patients with prehospital ECGs showing ventricular
tachycardia or ventricular fibrillation (i.e., these patients are managed by ACLS algorithms). We also
removed patients with confirmed prehospital STEMI, which included machine-generated ***ACUTE
MI*** warning, EMS-documentation of STEMI, and medical consult for potential CATH lab activation.

Independent reviewers extracted data elements from hospital systems on all patients meeting eligibility criteria. If a prehospital ECG had no patient identifiers, we used a probabilistic matching approach to link each encounter with the correct hospital record. This previously validated data linkage protocol was based on the ECG-stamped birth date, sex, and date/time logs, as well as based on EMS dispatch logs and receiving hospital records. All probabilistic matches were manually reviewed by research specialists for accuracy. The match success rate ranged from 98.6% to 99.8%.

403 Clinical Outcomes

404 Adjudications were made by independent reviewers at each local site after reviewing all available medical records within 30 days of the indexed encounter. Reviewers were blinded from all 405 ECG analyses and models' predictions. OMI was defined as coronary angiographic evidence of an 406 acute culprit lesion in at least one of the three main coronary arteries (LAD, LCX, RCA) or their primary 407 branches with TIMI flow grade of 0-1. TIMI flow grade of 2 with significant coronary narrowing > 70% 408 409 and peak troponin of 5-10.0 ng/mL was also considered indicative of OMI.^{17,21} These adjudications 410 were made by two independent reviewers. The Kappa coefficient statistic between the two reviewers was 0.771 (i.e., substantial agreement). All disagreements were resolved by a third reviewer. 411

ACS was defined per the fourth universal definition of myocardial infarction as the presence of symptoms of ischemia (i.e. diffuse discomfort in the chest, upper extremity, jaw, or epigastric area for more than 20 minutes) and at least one of the following criteria: (1) subsequent development of labile, ischemic ECG changes (e.g., ST changes, T inversion) during hospitalization; (2) elevation of cardiac troponin (i.e., > 99th percentile) during the hospital stay with rise and/or drop on serial testing; (3)

417	coronary angiography demonstrating greater than 70% stenosis, with or without treatment; and/or (4)
418	functional cardiac evaluation (stress testing) that demonstrates ECG, echocardiographic, or
419	radionuclide evidence of focal cardiac ischemia. ⁵ Patients with type 2 MI and pre-existing subacute
420	coronary occlusion were labeled as negative for ACS and OMI. This included around 10% of patients
421	with positive troponin but with no rise and/or drop in concentration on serial testing (i.e., chronic leak) or
422	with troponin leak attributed to non-coronary occlusive conditions such as pericarditis. On a randomly
423	selected small subset of patients (n=1,209), the Kappa coefficient statistic for ACS adjudication ranged
424	from 0.846 to 0.916 (i.e., substantial to perfect agreement).

425 ECG Methods

Prehospital ECGs were obtained in the field by paramedics as part of routine care. ECGs were 426 acquired using either Heart Start MRX (Philips Healthcare) or LIFEPAK-15 (Physio-Control Inc.) 427 monitor-defibrillator devices. All digital 12-lead ECGs were acquired at a sampling rate of 500 s/s (0.05-428 150 Hz) and transmitted to the respective EMS agency and receiving hospital. Digital ECG files were 429 exported in XML format and stored in a secondary server at each local site. ECG images were de-430 identified and manually annotated by independent reviewers or research specialists; ECGs with poor 431 quality or missing leads were removed from the study. Next, digital XML files were transmitted to the 432 433 Philips Advanced Algorithm Research Center for offline analysis (Cambridge, Massachusetts, USA). ECG featurization was described in detail elsewhere.¹⁸ Briefly, ECG signal preprocessing and 434 feature extraction were performed using a manufacturer-specific software (Philips DXL diagnostic 12/16 435 lead ECG analysis program). ECG signals were first preprocessed to remove noise, artifacts, and 436 437 baseline wander. Ectopic beats were removed, and median beats were calculated for each lead. Next, we used the root mean square (RMS) signal to identify global waveform fiducials, including the onset, 438 439 offset, and peak of the P wave, QRS complex, and T wave. Lead-specific fiducials were then identified to further segment individual waveforms into Q, R, R`, S, S`, and J point. 440

441	We then computed a total of 554 ECG features based on (1) the amplitude, duration, area,
442	slope and/or concavity of global and lead-specific waveforms; (2) the QRS and T axes and angles in
443	the frontal, horizontal, spatial, XY, XZ, and YZ planes, including directions at peak, inflection point, and
444	initial / terminal loops; (3) eigenvalues of the principal components of orthogonal ECG leads (I, II, V1-
445	V6), including PCA ratios for individual ECG waveform segments; and (4) T loop morphology
446	descriptors. Features with zero distribution were removed to prevent representation bias.
447	Next, we previously identified an optimal parsimonious list of the most important ECG features
448	that are mechanistically linked to cardiac ischemia as described in detail elsewhere. ¹⁸ Briefly, to prevent
449	omitted-feature bias, we used a hybrid approach that combines domain knowledge with a data-driven
450	strategy. First, clinical scientists identified 24 classical features that are known to correlate with cardiac
451	ischemia (i.e., lead-specific ST-80 and T wave amplitudes). Next, starting with a comprehensive list of
452	554 candidate features, we used data-driven algorithms (e.g., recursive feature elimination and
453	LASSO) to identify 198 supplemental features potentially related to ischemia. LASSO selects features
454	with non-zero coefficients after L1 norm regularization, and recursive feature elimination uses repeated
455	regression iterations to identify the features that have significant impact on model predictions. We then
456	examined the feature pairs in this expanded list of 222 features and removed features with very high

458 <u>selected by the model). Finally, we used feature importance ranking to identify the most parsimonious</u>

collinearity scores that contains redundant information (e.g., we kept QTc if both QT and QTc were

459 subset of features that are complementary and can boost the classification performance. This hybrid

460 approach eventually yielded a subset of 73 features that can serve as plausible markers of

457

461 <u>ischemiaClinical scientists initially reviewed a list of 554 features and marked the ones that are known</u>

462 to corrolate with cardiac ischemia. This list was then expanded by supplemental features identified by

463 data-driven algorithms (e.g., recursive feature elimination and LASSO). The clinical scientists then

464 reviewed the expanded list to examine feature pairs with high collinearity and retained the subset of

465	features that are complementary and can serve as plausible markers of ischemia. This approach
466	eventually yielded a subset of 73 features that was shown to boost classification performance. ¹⁸
467	Machine Learning Methods
468	We followed best practices recommended by "ROBUST-ML" and "ECG-AI stress test" checklists
469	to design and benchmark our machine learning algorithms. ^{51,59} To prevent measurement bias, ECG
470	features were manually reviewed to identify erroneous calculations. Physiologically plausible outliers
471	were replaced with ± 3 SD. On average, each feature had a 0.34% missingness rate (range 0.1% to
472	1.6%). Thus, we imputed missing values with the mean, median, or mode of that feature after
473	consultation with clinical experts. ECG metrics were then z-score normalized and used as input
474	features in machine learning models. The derivation and validation datasets were cleaned
475	independently to prevent data leakage. Both cohorts were recruited over the same time window,
476	suggesting the lack of temporal bias. To prevent potential mismatch with intended use, input features
477	for model development included only ECG data plus the machine-stamped age. No other clinical data
478	were used for model building.
479	We randomly split the derivation cohort into an 80% training set and a 20% internal testing set.
480	On the training set, we fit 10 machine learning classifiers: regularized logistic regression, linear
481	discriminant analysis, support vector machine, Gaussian Naïve Bayes, random forest, gradient
482	boosting machine, extreme gradient boosting, stochastic gradient descent logistic regression, k-nearest
483	neighbors, and artificial neural networks. Each classifier was optimized over 10-fold cross validation to
484	finetune hyperparameters. After selecting optimal hyperparameters, models were re-trained on the
485	entire training subset to derive final weights and create a lockout model to evaluate on the holdout test
486	set. We calibrated our classifiers to produce a probabilistic output which can be interpreted as a
487	confidence level (probability risk score). Trained models were compared using the AUROC curve with
488	Wilcoxon signed-rank test for pairwise comparisons. ROC-optimized cutoffs were chosen using Youden
489	index, and classifications on confusion matrix were compared using McNemar's test.

490 The random forest classifier (RF) achieved high accuracy on the training set (low bias) with a 491 relatively small drop in performance on the test set (low variance), indicating an acceptable bias-492 variance tradeoff and low risk of overfitting (Extended Data Fig. 78). Although the support vector 493 machine (SVM) model had lower variance on the test set, when compared with the RF model, there 494 were no significant differences in AUROC (Delong's test) or their binary classifications (McNemar's 495 test). Moreover, there were no differences between the RF and SVM models in terms of Kolmogorov-Smirnov goodness-of-fit (0.716 vs. 0.715) or the Gini purity index (0.82 vs. 0.85). Due to its scalability 496 and intuitive architecture, we chose the probability output of the RF model to build our derived OMI 497 score. We generated density plots of these probability scores for positive and negative classes and 498 499 selected classification thresholds for low, intermediate, and high-risk groups based on prespecified 500 NPV > 0.99 and TPR > 0.50. Finally, we used the lock-out random forest classifier to generate 501 probability scores and risk classes on the completely unseen external validation cohort. The code to 502 generate probability scores is included with the supplemental materials of this manuscript.

503 Reference Standard

To reduce the risk of evaluation bias, we benchmarked our machine learning models against 504 505 multiple reference standards used during routine care in clinical practice. First, we used a commercial, FDA-approved, ECG interpretation software (Philips DXL diagnostic algorithm) to denote the likelihood 506 of ischemic myocardial injury. This likelihood (yes/no) was based on a composite of the followings: (1) 507 diagnostic codes for ">>>Acute MI<<<", including descriptive statements that denote "acute", "recent", 508 509 "age indeterminate", "possible" or "probable"; and (2) diagnostic codes for ">>>Acute Ischemia<<<", including descriptive statements that denote "possible", "probable", or "consider". Diagnostic statements 510 that denoted "old" [infarct], "nonspecific" [ST depression], or "secondary to" [LVH or high heart rate] 511 512 were excluded from this composite reference standard.

513 We also used practicing clinicians' overread of ECGs to denote the likelihood of ischemic 514 myocardial injury <u>on a given ECG (yes/no) when a STEMI pattern does not exist, which is congruent</u>

515	with how ED physicians evaluate these patients in clinical practice. Independent physician reviewers
516	annotated each 12-lead ECG image as per the fourth universal definition of MI criteria, ⁵ including two
517	contiguous leads with ST-elevation (≥ 0.2 mV for V2-V3 in men ≥ 40 years and ≥ 2.5 mm in men < 40
518	years; \geq 0.15 mV for V2-V3 in women; or \geq 0.1 mV in other leads) or ST-depression (new horizontal or
519	down-sloping depression \ge 0.05 mV); with or without T wave inversion (> 0.1 mV in leads with
520	prominent R wave or R/S ratio > 1). Reviewers were also prompted to use their clinical judgment to
521	identify highly suspicious ischemic changes (e.g., reciprocal changes, hyperacute T waves), as well as
522	to account for potential confounders (e.g., bundle branch blocks, early repolarization). On a randomly
523	selected subset of patients in the derivation cohort (n=1,646), the Kappa coefficient statistic between
524	two emergency physicians who interpreted the ECGs was 0.568 (i.e., moderate agreement). A third
525	reviewer was used to adjudicate discrepancies on this randomly selected subset. Similarly, on a
526	randomly selected subset of patients in the external validation cohort (n=375), the Kappa coefficient
527	statistic between the two board-certified cardiologists who interpreted the ECGs was 0.690 (i.e.,
528	substantial agreement).
529	Finally, given that clinicians largely depend on risk scores to triage patients in the absence of STEMI,
530	which would significantly affect how OMI patients are diagnosed and treated in clinical practice, we
531	compared our derived OMI risk score against the HEART risk score. This score is commonly used in
532	US hospitals and it has been well-validated for triaging patients in the emergency department. ⁶⁰ The
533	HEART score is based on the patient's <u>H</u> istory at presentation, <u>E</u> CG interpretation, <u>Age</u> , <u>R</u> isk factors,
534	and initial <u>T</u> roponin values (range 0-10). This score places patients in low (0-3), intermediate (4-6), and
535	high-risk (7-10) groups. Given that troponin results are not usually available at first medical contact, we
536	used a modified HEAR score after dropping the ${f T}$ roponin values, which has also been previously

- 537 validated for use by paramedics prior to hospital arrival.³⁶ The comparison against the HEART score
- 538 herein focused on establishing the incremental gain of using the derived OMI score over routine care at
- 539 initial triage. We compared how the new risk classes assigned by our derived OMI score

agree with or differ from the risk classes assigned by the HEART score, which could inform potential incremental gain over routine care.

542 Statistical_Analysis

Descriptive statistics were reported as mean ± standard deviation or n (%). Missing data was assessed for randomness and was handled during ECG feature selection (see Machine Learning Methods section above). Normality of distribution was assessed prior to hypothesis testing where deemed necessary. ECG features were z-score normalized as part of standard input architectures for machine learning models. Comparisons between cohorts were performed using chi-square (for discrete variables) and independent samples t-test or Mann-Whitney U test (for continuous variables). The level of significance was set at alpha 0.05 for two-tailed hypothesis testing where applicable.

All diagnostic accuracy values were reported as per STARD recommendations (Reporting 550 Guidelines for Diagnostic Accuracy Studies). We reported classification performance using AUROC 551 curve, sensitivity (recall), specificity, PPV (precision), and NPV, along with 95% confidence interval (CI) 552 where applicable. For 10-fold cross validation, we compared the multiple classifiers using the Wilcoxon 553 signed-rank test (for AUROC curves) and McNemar's test (for confusion matrices). We derived low-, 554 intermediate-, and high-risk categories for the final classifier using Kernel density plot estimates 555 between classes. The adequacy of these risk classes was evaluated using Log-rank chi-square of 556 557 accumulative risk for clinically important outcomes over the length of stay during the indexed 558 admission.

559 For assessing the incremental gain in classification performance, we compared the AUROC of 560 the final model against reference standards using DeLong's test. For ease of comparison, the 561 confidence bounds for AUROC of the reference standards (commercial system and practicing 562 clinicians) were generated using 1000 bootstrap samples. <u>To place the incremental gain value in a</u> 563 <u>broader context of the clinical workflow, We we then also</u> computed the Net Reclassification

564 Improvement (NRI) index of our model against the HEAR score during the initial assessment at first

- 565 medical contact. Risk scores are an integral part of clinical workflow in patients with suspected ACS
- 566 who do not meet STEMI criteria. As per STARD recommendations (Reporting Guidelines for Diagnostic
- 567 Accuracy Studies), the NRI Index evaluates the net gain between up-triage and down-triage when
- 568 correctly reclassifying risk class assignments of an "old" test (HEART score) using a "new" test (the
- 569 derived OMI score).
- . 570 We used logistic regression to identify the independent predictive value of OMI risk classes. We
- 571 used variables significant in univariate analysis and then built multivariate models with stepwise
- 572 backward selection method using Wald chi-square criteria. We reported odds ratios with 95% CI for all
- significant predictors. All analyses were completed using Python v3.8.5 and SPSS v24.

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1		

Table 1. Baseline demographic and clinical characteristics 749

	DERIVATION & TESTING COHORT (N=4,026)	EXTERNAL VALIDATION COHORT (N=3,287)
AGE (YEARS)	59±16 (18-102)	60±15 (21-100)
SEX Male Female	2,122 (53%) 1,904 (47%)	1,814 (55%) 1,473 (45%)
RACE White Black Others Unknown	1,698 (42%) 1,328 (33%) 52 (1.3%) 948 (24%)	1,326 (40%) 1,544 (47%) 40 (1%) 377 (12%)
ETHNICITY Not Hispanic Hispanic / Latino Unknown	3,043 (76%) 19 (1%) 964 (23%)	2,850 (87%) 116 (3.5%) 321 (9.5%)
PAST MEDICAL HISTORY Hypertension Diabetes High cholesterol Current smoker Known CAD Prior MI Prior PCI Prior CABG		2,090 (64%) 1,067 (33%) 1,376 (42%) 802 (25%) 964 (30%) 929 (29%) 134 (4%) 470 (14%)
ECG & LAB FINDINGS Sinus rhythm Atrial fibrillation Left BBB Right BBB ECG-LVH cTnl positive (serial testing) MEDICAL THERAPY	3,496 (87%) 354 (9%) 94 (2.3%) 237 (5.9%) 383 (9.5%) 330 (8.2%) 729 (18.1%)	2,614 (80%) 352 (11%) 114 (3.5%) 215 (6.6%) 467 (14.2%) 736 (22.4%) 1,177 (35.8%)
PCI (ANY STENT) Emergent PCI (<90 MIN) Total LAD occlusion Total LCX occlusion Total RCA occlusion CABG	300 (7.5%) 144 (3.6%) 91 (2.3%) 63 (1.6%) 101 (2.5%) 34 (0.8%)	245 (7.5%) 157 (4.8%) 94 (2.9%) 88 (2.7%) 102 (3.1%) 30 (0.9%)
STUDY OUTCOMES CONFIRMED ACS OMI Other Acute MI (NOMI) Unstable Angina 30-DAY CV DEATH	550 (13.7%) 210 (5.2%) 240 (6.0%) 100 (2.5%) 137 (3.4%)	537 (16.3%) 209 (6.4%) 220 (6.7%) 108 (3.3%) 111 (3.4%)

750 Values are mean ± SD (min-max) or n (%); CAD: coronary artery disease; MI: myocardial infarction; BBB: bundle

751

branch block; LVH: left ventricular hypertrophy; PCI: percutaneous coronary intervention; LAD: left anterior descending artery; LCX: left circumflex artery; RCA: right coronary artery; CABG: coronary artery bypass graft; 752

753 OMI: occlusion MI; NOMI: non-occlusion MI; CV: cardiovascular.

754 FIGURE LEGENDS

- 755 Fig. 1. Cohort and sample selection
- 756 This flow diagram shows patient inclusion and exclusion in each cohort, as well as the dataset partition
- 757 for training, internal testing, and external validation. Exclusions are not mutually exclusive.
- 758 Fig. 2. Algorithm derivation and testing
- 759 This figure shows (A) the classification performance of the machine learning model against other
- reference standards for detecting occlusion myocardial infarction (OMI), (B) the probability density plots
- of OMI(+) and OMI(-) classes as denoted by the machine learning model, along with optimal cutoffs of
- 762 low-risk, intermediate, and high-risk, and (C) distribution of patients in low-risk (+), intermediate risk
- 763 (++) and high-risk (+++) as per the machine learning model and HEART score.

764 Fig. 3. Model explainability for OMI detection

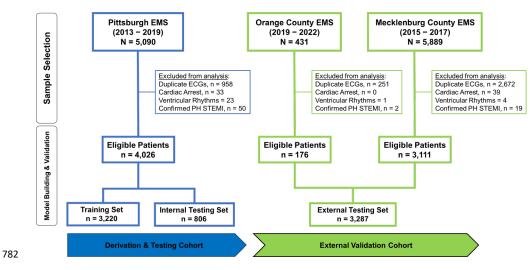
- 765 This figure shows (A) SHAP values for the 25 most important features driving the predictions of the
- 766 machine learning classifier in the derivation cohort, and (B) the aggregate median beats of ECGs with
- occlusion myocardial infarction (OMI) class (red) and the aggregate median beats of ECGs with normal
- sinus rhythm and no OMI (blue).

769 Fig. 4. External validation of ECG-SMART algorithm

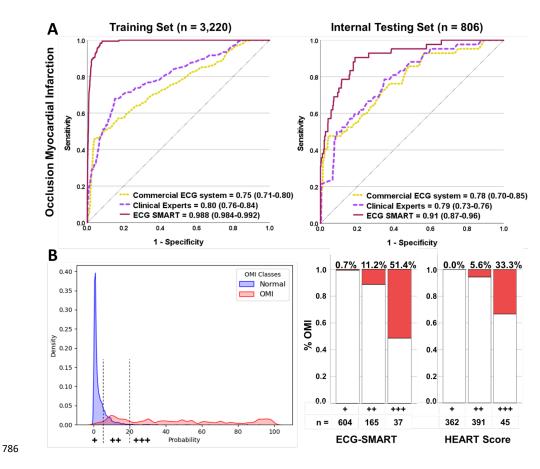
- 770 This figure shows (A) the classification performance of the machine learning model against other
- 771 reference standards for detecting occlusion myocardial infarction (OMI), (B) the independent clinical
- 772 predictors of OMI on multivariate logistic regression testing, and (C) the overall sensitivity and
- 573 specificity (95% confidence interval [CI]) of the derived OMI score, along with breakdown across
- 574 subgroups based on age, sex, comorbidities, and baseline ECG findings. The size of markers denotes
- the sample size of the respective subgroup.
- 776 Fig. 5. Net reclassification improvement of OMI risk score when integrated in the clinical
- 777 workflow at first medical contact Decision analysis for the incremental gain of OMI risk score in
- 778 reclassifying patients

- 779 This figure simulates the incremental gain of the derived risk score in reclassifying the initial triage
- 780 decisions by emergency personnel at first medical contact.

Fig. 1. Cohort and sample selection







788 Fig. 3. Model explainability for OMI detection

st80_III	
st80_aVL	• • • • • • • • • • • • • • • • • • •
tamp_aVL	• • • • • • • • • • • • • • • • • • •
STT_PCAratio	·
st80_V2	•
ТрТе	• • • • • • • • • • • • • • • • • • •
st80_aVF	
tamp_III	
st80_I	
pcaTamp	
st80_V6	
fpTaxis	
st80_V1	
fpTinfl1Axis	
PCA2	
TampInfl1	
Age	
tamp_aVF	
antConcaveAmp	
TCRT	
T_PCAratio	
st80_V4	·····
HR	
tamp_V4	
st80_V3	Low
	-0.05 0.00 0.05 0.10
	SHAP value (impact on model output)



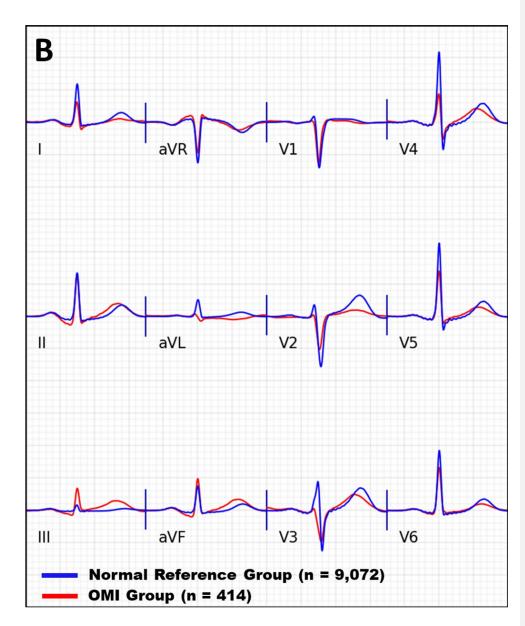
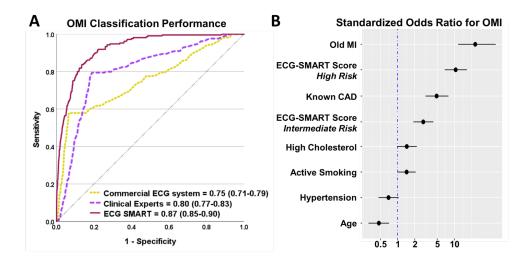


Fig. 4. External validation of ECG-SMART algorithm for OMI detection 792



Sensitivity (95% CI) ы

	N				
Age < 65-year-old	2,057			—	
Age ≥ 65-year-old	1,227				
Male Sex	1,814		-		
Female Sex	1,470			-	
White Race	1,336		-		
Black Race	1,544	-	-		
Hypertension	2,090		Ē		
No Hypertension	1,050		·	•	-
Diabetes	1,067		-	<u> </u>	
No Diabetes	2,220				
Known CAD	964				
No known CAD	2,323	۰		-	
Current smoker	802		,	-	
Not current smoker	2,309	F			
Sinus Rhythm	2,524		H	—	
Not sinus rhythm	762			<u> </u>	
Narrow QRS duration	2,485		—	—	
Wide QRS duration	626		F	-	
ECG-LVH	467			-	
No ECG-LVH	2,820		—		
Overall	0.86 (0	.81-0.91)			
		0.70	0.80	0.90	1.00

Male Sex 1,821 Female Sex 1,473 White Race 1,332 Black Race 1,547 Hypertension 2,097 No Hypertension 1,053 Diabetes 1,070 No Diabetes 2,227 Known CAD 973 No known CAD 2,324 Current smoker 807 Not current smoker 2,314 Sinus Rhythm 2,533 Not sinus rhythm 763 Narrow QRS duration 2,493 Wide QRS duration 628 ECG-LVH 467 No ECG-LVH 2,830 Overall 0.98 (0.97-0.99) +++ 0.90

Ν

2,063

1,231

Age < 65-year-old

Age ≥ 65-year-old

1.00

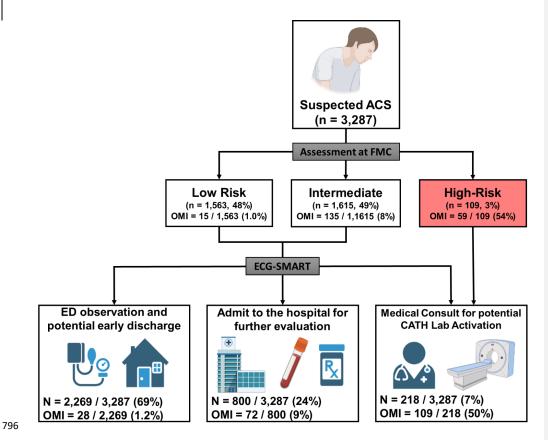
Specificity (95% CI)

793

С

794 Fig. 5. Decision analysis for the incremental gain Net reclassification improvement of OMI risk

795 score when integrated in the clinical workflow re-classifying patients at first medical contact



DATA SUPPLEMENT

798

799 Legend for Extended Data Figures

800 Extended Data Fig. 1. The relationship between the magnitude of vessel occlusion and the

801 classification of acute coronary events

802 This figure shows the spectrum of coronary artery disease (CAD) as a function of severity and extent of 803 atherosclerosis plaque progression, ranging from patent coronary artery (far left) to total coronary occlusion (far right). Among patients who develop symptomatic CAD, including those evaluated for 804 805 chest pain or angina-like symptoms, a subset is diagnosed with acute coronary syndrome (ACS). This group is subclassified as either acute myocardial infarction (MI) or unstable angina (UA). Those with 806 807 acute MI can be further subclassified, based on the presence of ST-elevation on the ECG, as either ST-808 elevation myocardial infarction (STEMI) or without ST-elevation (NSTEMI). The STEMI and NSTEMI patients overlap in terms of the presence or absence of total occlusion (depicted as triangles across the 809 810 continuum in the figure). Alternatively, the same group with acute MI can be subclassified, based on angiographic TIMI flow criteria, as either occlusion (OMI) or non-occlusion (non-OMI) myocardial 811 812 infarction. Unlike STEMI, OMI classification better aligns with focal angiographic findings since this 813 group exclusively contains patients with total coronary occlusion. The color gradient indicates the severity of disease. This Figure was created with BioRender.com. Reproduced with permission from Al-814 815 Zaiti et. al.¹ (permission number 5471421247333, Licensed content publisher: Elsevier). Extended Data Fig. 2. Graphical abstract summarizing the flow of study and main findings 816 817 This figure provides a graphical summary of the study flow and main findings.

818 Extended Data Fig. 3. Local explainability of feature importance on a selected example

- 819 This figure shows the Bbaseline ECG of a 50-year-old female with a past medical history of
- 820 hypertension, high cholesterol, prior myocardial infarction, and current smoking. The ECG was
- 821 documented as benign with isolated non-specific T wave changes, and the patient was triaged as

822	intermediate risk. The OMI score was 62 indicating the need to up-triage. The patient was later sent to
823	the catheterization lab where she had complete occlusion of the right coronary artery. The OMI score
824	on this baseline ECG was 62 indicating high risk designation. The force plot identified the five most
825	important ECG features that met the contribution threshold of the random forest model: negative T
826	wave in aVL, slight ST depression in aVL and V2, and slight ST elevation in aVF and III.
827	Extended Data Fig. <u>34</u> . Selected examples of <u>a patients correctly reclassified as OMI</u>
828	This figure shows two examples of <u>an ECG that was patients who were</u> correctly reclassified as
829	occlusion myocardial infarction by the machine learning model. (A)- <u>This Baseline-baseline</u> ECG was for
830	of a 67-year-old male with a past medical history of high cholesterol and a prior myocardial infarction.
831	The ST-depression in anterior-lateral leads were noted, and the patient was triaged as intermediate
832	risk. The OMI score was 49 indicating the need to up-triage. The patient was later sent to the
833	catheterization lab where he had complete occlusion of the right coronary artery. (B) Baseline ECC of a
834	50 year old female with a past medical history of hypertension, high cholesterel, prior myecardial
835	infarction, and current smoking. The ECG was documented as benign with isolated non specific T wave
836	changes, and the patient was triaged as intermediate risk. The OMI score was 62 indicating the need to
837	up triage. The patient was later cent to the catheterization lab where che had complete occlusion of the
838	right coronary artory.
839	Extended Data Fig. 4 <u>5</u> . Selected example of a missed OMI by our model
840	This figure provides a selected example of a patient with occlusion myocardial infarction that was
841	missed by the machine learning model and other reference standards. This ECG was obtained on a 70-
842	year-old female with a past medical history of hypertension, high cholesterol, prior myocardial
843	infarction, and current smoking. The baseline clinical interpretation suggests normal sinus rhythm with
844	benign findings. There are isolated Q waves in inferior leads, low ECG voltage, and some baseline

- 845 wander and high frequency noise in few leads. The OMI risk score was 2 indicating a low risk. The
- 846 patient was later sent to the catheterization lab, which showed significant left main occlusion and had

- 847 many stents placed. The patient developed new-onset HF during hospitalization. A closer look at this
- 848 ECG by experienced ECG readers suggests that this ECG could resemble the "precordial swirl
- 849 pattern", a rightward ST-elevation vector, with STE in V1 and aVR and reciprocal ST-depression in V5
- and V6. This pattern was found to correlate with LAD occlusion.

851 Extended Data Fig. <u>56</u>. Development and validation of an algorithm to screen for any ACS event

- 852 This figure shows the classification performance of the machine learning model against other reference
- standards for detecting any acute coronary syndrome event (ACS). The figure also shows the
- distribution of patients in low-risk, intermediate risk, and high-risk groups as per our derived risk score.
- 855 There is a notable gain in precision (rule-in) but a significant loss in recall (rule-out).

856 Extended Data Fig. 67. Limitations of ST amplitude on surface ECG as a sole marker of

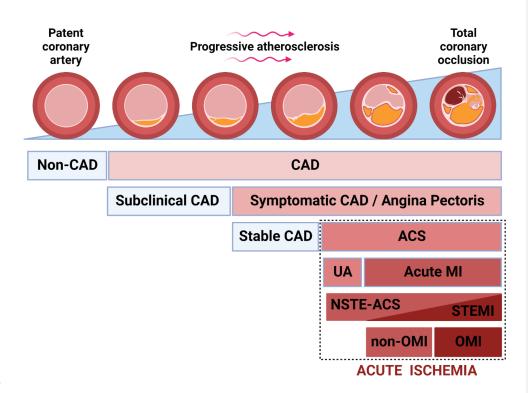
- 857 myocardial ischemia
- 858 This figure shows: (A) cardiac model of anterior wall epicardial ischemia with corresponding ST-
- elevation on V3 to V5 of the 12-lead ECG. (B) cardiac model of anterolateral and inferior-apical
- 860 epicardial ischemia with corresponding attenuation of ST changes on the 12-lead ECG. This figure was
- 861 generated using ECGSIM (<u>www.ecgsim.org</u>). Reproduced with permission from Al-Zaiti et. al.¹
- 862 (permission number 5471421247333, Licensed content publisher: Elsevier).

Extended Data Fig. 78. Comparison between 10 algorithms trained on the derivation cohort to
classify OMI

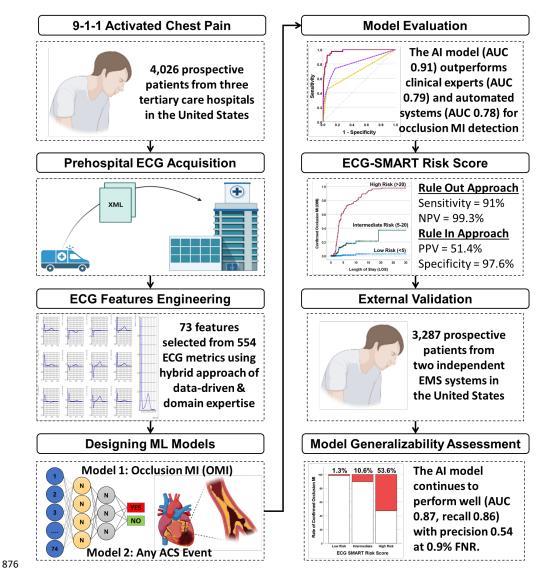
- 865 This figure compares the area under the receiver operator characteristics curves (95% confidence
- interval) of 10 classifiers during training (left) and testing (right) on the derivation cohort. RF: random
- 867 forest; KNN: K-nearest neighbors; GBM: gradient boosting machine; XGB: extreme gradient boosting;
- 868 SVM: support vector machine; ANN: artificial neural networks; LogReg: regularized logistic regression;
- 869 LDA: linear discriminant analysis; SGD_LogReg: stochastic gradient descent logistic regression; G_NB:
- 870 Gaussian Naïve Bayes.

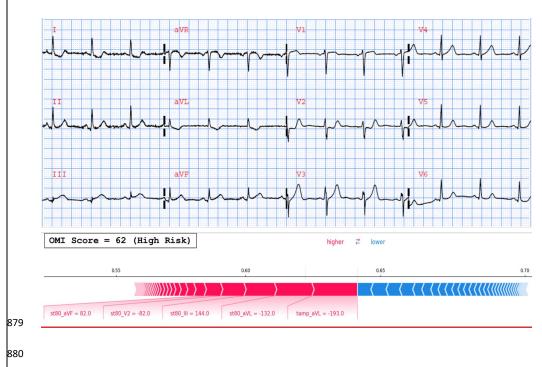
872 Extended Data Fig 1. The relationship between magnitude of coronary occlusion and coronary

873 artery disease and acute events classification



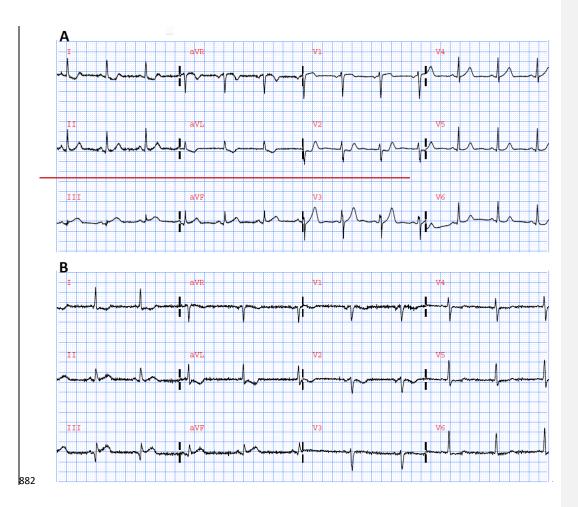
875 Extended Data Fig 2. Graphical abstract summarizing the flow of study and main findings

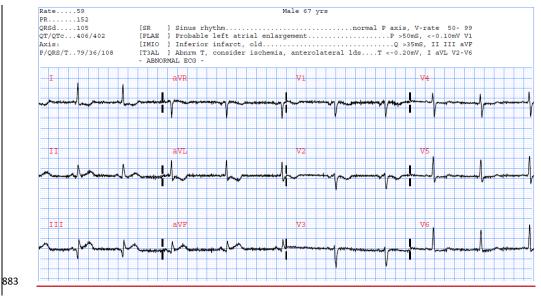




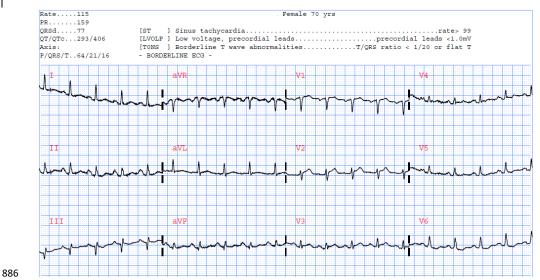
878 Extended Data Fig. 3. Local explainability of feature importance on a selected example

881 Extended Data Fig <u>34</u>. Selected examples of <u>a patients</u> correctly reclassified as OMI



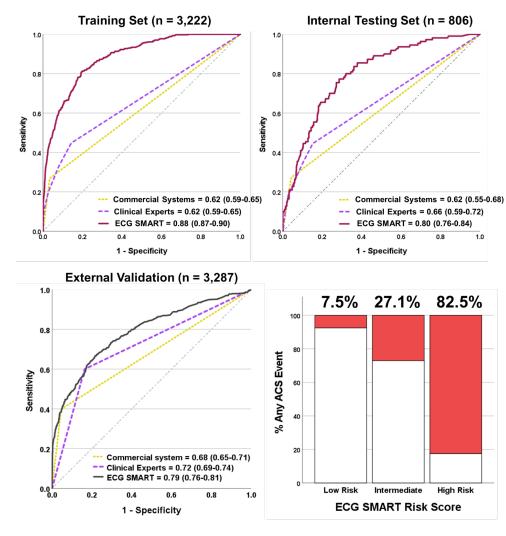






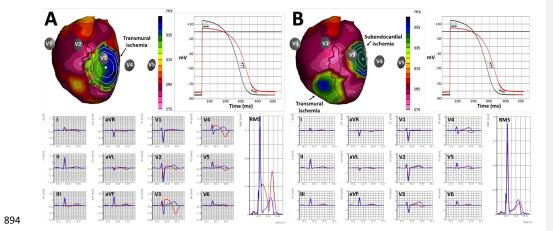
885 Extended Data Fig 45. Selected example of a missed OMI by our model





892 Extended Data Fig. 67: Limitations of ST amplitude on surface ECG as a sole marker of

893 myocardial ischemia



895

897 Extended Data Fig 78. Comparison between 10 algorithms trained on the derivation cohort to

898 classify OMI

