

An investigation of stride interval stationarity while listening to music or viewing television

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Abstract

In recent years, there has been considerable interest in the effects of auditory and visual distractions on pedestrian ambulation. A fundamental temporal characteristic of ambulation is the temporal fluctuation of the stride interval. In this paper, we investigate the stationarity of stride interval time series when people are exposed to different forms of auditory and visual distractions. An increase in nonstationary behavior may be suggestive of divided attention and more frequent central modulation of locomotion, both of which may have ramifications on pedestrian vigilance and responsiveness to environmental perturbations. One group of fifteen able-bodied (6 females) young adult participants completed a music protocol (overground walking with and

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without music). A second group of fifteen (7 females) did a television protocol (treadmill walking while watching TV with and without sound). Three walking trials, each 15 minutes in duration, were performed at each participant's comfortable walking speed, with force sensitive resistors under the heel of each foot. Using the reverse arrangements test, the vast majority of time series were nonstationary, with a time-varying mean as the principal source of nonstationarity. Furthermore, the television trial with sound had the greatest number of nonstationarities followed by overground walking while listening to music. We discuss the possibility that these conditions measurably affect gait dynamics through a subconscious synchronization to external rhythms or a cyclic distraction followed by a period of increased conscious correction of gait timing. Our findings suggest that the regulation of stride timing is particularly susceptible to constant, time-evolving auditory stimuli, but that normal pacing can be restored quickly upon stimulus withdrawal. These kinds of sensory distractions should thus be carefully considered in studies of pedestrian ambulation.

Key words: Gait, stride intervals, stationarity, music, television, fractal analysis.

1. Introduction

Walking is the most common physical activity among adults (Owen et al., 2004). One of the fundamental temporal characteristics of walking is the stride interval, that is, the time between successive heel strikes. Walking is a complicated task governed by the hierarchical control of the motor cortex, spinal pattern generators and feedback from the vestibular system (Di-

7 etz, 2002; West and Griffin, 1999). Due to these varying levels of conscious
8 and autonomic control, interstride times fluctuate in a very complex manner
9 (West and Griffin, 1999; Hausdorff et al., 1999, 2001). These temporal vari-
10 ations of stride intervals have been shown to differentiate between persons
11 with and without pathologies, young and older adult gait, and overground
12 and treadmill walking (Hausdorff et al., 1997; Chau and Rizvi, 2002).

13 The stationarity of these complex fluctuations may reveal the magnitude
14 with and time scale over which the nervous system makes adjustments as
15 a person walks. Using recurrence plots, a treadmill study reported station-
16 ary joint angles and body accelerations while walking and attributed this
17 temporal invariance to the constant speed imposed on the participant (Ding-
18 well and Cusumano, 2000). However, a subsequent pediatric treadmill study
19 found nonstationarities via the reverse arrangement test in stride interval
20 time series despite speed constraints (Fairley et al., 2010), implicating the
21 children’s evolving comfort level with the treadmill and varying attentiveness
22 to the task as possible sources of nonstationarity. The temporal variability
23 of stride intervals may thus be connected to the physical and cognitive ca-
24 pability to dynamically adapt one’s gait and more generally, to respond to
25 changes in the environment.

26 This study aimed to determine whether or not stride interval time series
27 remain stationary when subjects are exposed to different audio and visual
28 stimuli. In particular, we considered the stationarity of the stride interval
29 time series while subjects walked overground when listening to music or while
30 they walked on a treadmill and watched television. The effects of these
31 sensory stimuli on gait are of concern given the increasing pedestrian use of

32 personal music players (Bungum et al., 2005), handheld gaming devices, text
33 messaging applications, portable video players, and pedestrian navigation
34 systems (Torres-Solis and Chau, 2010). Additionally, video and music are
35 becoming important tools in rehabilitation research and practice (Sveistrup,
36 2004; Fung et al., 2006; Schauer and Mauritz, 2003).

37 **2. Methodology**

38 *2.1. Participants*

39 Two groups of participants were recruited. The first group consisted of 15
40 able-bodied participants (6 females; mean age 25.9 ± 2.8 years). The second
41 group consisted of 15 able-bodied participants (7 females; mean age $25.4 \pm$
42 2.7 years). All participants were recruited from the Bloorview Research In-
43 stitute and provided written informed consent in accordance with Bloorview
44 Research Ethics Board. Subjects met the inclusion criteria of having normal
45 or corrected-to-normal vision, intact hearing and right-foot dominance. Par-
46 ticipants were excluded if they had any history of neurological pathology that
47 would have compromised natural bipedal ambulation. For both protocols,
48 participants were instructed to wear comfortable walking shoes.

49 *2.2. Data collection*

50 Data were collected using two different protocols: music and television.
51 The first group completed the music protocol, which comprised two iden-
52 tical sessions, separated by at least 24 hours, and each consisting of three
53 15 minute trials in the following sequence: 1. overground walking (OW-
54 NoMusic1), 2. overground walking with music (OW-Music), and 3. over-

55 ground again (OW-NoMusic2). Unconstrained overground walking was per-
56 formed at a self-selected pace on an indoor rectangular path (width ~ 2.5 m,
57 length ~ 100 m) in sparsely populated hallways with linoleum flooring. Par-
58 ticipants were instructed to “walk at a comfortable pace”. An investigator
59 walked slightly behind the subject during the test to record any stumbles or
60 falls which might skew the stride time distribution. Songs were selected from
61 a “Top 40s” list for that month and remained consistent for each participant.
62 However, the order of song presentation was randomized for each participant.

63 The second group completed the television protocol, which comprised two
64 identical sessions (at least 24 hours apart) consisting of three 15 minute trials
65 each. These trials invoked the following conditions in randomized order: 1.
66 walking on a treadmill while viewing television (TW-TV), 2. walking on
67 a treadmill while viewing television without sound (TW-TVNoSound), and
68 3. walking on a treadmill while viewing television with subtitles (no sound)
69 (TW-TVSubtitles). Television programs were presented in random order
70 from a DVD of popular movie “shorts”. Hence, each participant watched
71 the same videos but in different order. At the start of each session, the
72 participant completed a 5 minute warm-up on the treadmill at a self-selected,
73 comfortable speed. Prior to each trial, the participant also did a 2 minute
74 warm-up at his or her preferred speed, subject to the same TV condition
75 (sound, no sound or subtitles) as the corresponding trial. Subsequent to the
76 warm-up, the participant rested for 45 seconds (quiet standing) and then
77 began walking on the treadmill at his or her preferred speed. The preferred
78 speed was determined prior to the trial with the participant walking slowly
79 on the treadmill. The speed was increased in 0.1 mph increments until the

80 participant reported that his or her preferred speed had been reached. The
81 speed was then increased by 0.5 mph and then sequentially decremented by
82 0.1 mph until the participant reported attaining his or her preferred speed.
83 The mean speed obtained by the above procedure was taken as the preferred
84 walking speed. A motorized treadmill (GaitKeeper; Mobility Research) was
85 used for the treadmill conditions.

86 Force-sensitive resistors (Model no. 406, Interlink Electronics) under the
87 insole of each shoe of the subject were used to measure heel strikes. A change
88 in voltage indicating a heel contact with the walking surface was sampled at
89 200 Hz. These data were recorded by a custom-made datalogger, constructed
90 by mounting a programmable processor (R-Engine-A, Tern Inc.) inside an
91 enclosure. The datalogger was carried by the participant in a waist pouch.
92 The FSRs fed into the datalogger via two wires along the length of the lateral
93 side of each leg. The examiner ensured that the wires and waist pouch did
94 not impede natural gait.

95 *2.3. Stationarity and reverse arrangements test*

96 Suppose $x(n)$ represents observations made during the time interval $0 \leq$
97 $n \leq N - 1$, where N represents the length of the signal. From the time series
98 point of view, we can consider $x(n)$ to be a discrete-time series since the
99 observations are made at time intervals from the discrete set Υ of times (with
100 N representing the cardinality of Υ). The time series can then be considered
101 a realization of the family of real-valued random variables $\{\chi_n, n \in \Upsilon\}$ that
102 constitute a stochastic process defined on a probability space (Brockwell and
103 Davis., 1991).

104 Stationarity is a property of a time series in which the probability dis-

105 tribution of values of the series are independent of time translations. In
 106 other words, a time series is strictly stationary if the cumulative distribution
 107 function of the joint distribution, $F_{\chi_{n_1}, \dots, \chi_{n_N}}(x_1, x_2, \dots, x_N)$, is invariant to
 108 a shift in the origin, i.e.,

$$F_{\chi_{n_1}, \dots, \chi_{n_N}}(x_1, x_2, \dots, x_N) = F_{\chi_{n_1+\tau}, \dots, \chi_{n_N+\tau}}(x_1, x_2, \dots, x_N) \quad (1)$$

109 for all positive τ . This is also referred to as strong stationarity as opposed to
 110 weak or wide-sense stationarity, in which only the first two moments of the
 111 series are required to be time-invariant:

$$E\{\chi_{n_1}\} = E\{\chi_{n_1+\tau}\} \quad (2)$$

$$\text{Cov}(\chi_{n_1}, \chi_{n_2}) = \text{Cov}(\chi_{n_1+\tau}, \chi_{n_2+\tau}). \quad (3)$$

113 The non-parametric reverse arrangement test (RAT) has often been used to
 114 test the wide-sense stationarity of a time series (Bendat and Piersol, 2000;
 115 Alves and Chau, 2008; Chau et al., 2005; Cao et al., 1997). The test searches
 116 for monotonic trends in the mean square values calculated within nonover-
 117 lapping intervals of a particular signal of interest. The steps of the reverse
 118 arrangement test are as follows (Bendat and Piersol, 2000; Chau et al., 2005):

- 119 1. Divide the time series into K nonoverlapping segments, with the as-
 120 sumption that the data within each segment is independent. If a priori
 121 knowledge about the length of a segment exists, then the number of
 122 segments can be calculated as:

$$K = \left\lfloor \frac{N}{L} \right\rfloor \quad (4)$$

123 where L is the desired segment length and $\lfloor \bullet \rfloor$ represents the great-
 124 est integer function. It is clear that N is not necessarily an integer

125 multiple of K . Hence, some of the data points must be omitted. Pre-
 126 vious research on stride intervals time series showed that there are no
 127 significant statistical differences between stationarity estimates from
 128 different data trimming approaches (Fairley et al., 2010). Therefore,
 129 the data were trimmed from both sides as suggested in Fairley et al.
 130 (2010), and the trimmed version of the signal is denoted by $x_t(m)$ where
 131 $0 \leq m \leq M - 1$ and $M = KL \leq N$.

132 2. Form a vector $y \in \mathbb{R}^K$ whose points are assigned as follows:

$$y(k) = \frac{1}{L} \sum_{j=kL}^{(k+1)L-1} x_t^2(j) \quad \text{for } 0 \leq k \leq K - 1 \quad (5)$$

133 3. A reverse arrangement occurs when $y(a) > y(b)$ for $a < b$. Hence, using
 134 this simple rule, for $y(k)$ form an indicator, $i(k, d)$ as follows:

$$i(k, d) = \begin{cases} 1 & \text{if } y(k) > y(k + d) \\ 0 & \text{otherwise} \end{cases} \quad (6)$$

135 for $1 < d \leq D$ where $D = K - k - 1$. Therefore, the number of reverse
 136 arrangements for the k^{th} time step is given by

$$I(k) = \sum_{d=1}^D i(k, d) \quad (7)$$

137 and the total number of reverse arrangements is given by

$$I_T = \sum_{k=0}^{K-1} I(k). \quad (8)$$

138 4. For a stationary process, the distribution of I_T is approximately normal
 139 and its expected value is given by

$$\mu_T = \frac{L(L - 1)}{4} \quad (9)$$

140 and its variance by

$$\sigma_T^2 = \frac{L(L-1)(2L-5)}{72} \quad (10)$$

141 Therefore, the null hypothesis is that I_T comes from a normal distri-
142 bution with its mean and variance given by (9) and (10), respectively.

143 The null hypothesis is rejected at a significance level α if I_T falls outside
144 the corresponding critical values.

145 In this paper, the test statistic defined as

$$z_T = \frac{I_T - \mu_T}{\sigma_T} \quad (11)$$

146 is used, with the assumption that $z_T \sim \mathcal{N}(0, 1)$. The critical values at
147 significance level α are then $z_{1-\alpha/2}$ and $z_{\alpha/2}$ where z is a standard normal
148 variate, and for a 5% significance level these are given by $z_{\alpha/2} = -1.96$ and
149 $z_{1-\alpha/2} = 1.96$. The values of the test statistics, z_T , can fall within one of the
150 three possibilities:

- 151 • $z_T \leq z_{\alpha/2} \rightarrow$ There are fewer reverse arrangements than expected of
152 a stationary signal, implying the presence of an upward trend in the
153 mean square sequence.
- 154 • $z_T \geq z_{1-\alpha/2} \rightarrow$ There are more reverse arrangements than expected of
155 a stationary signal, implying the presence of a downward trend in the
156 mean square sequence.
- 157 • $z_{\alpha/2} < z_T < z_{1-\alpha/2} \rightarrow$ The null hypothesis that a time series is (weakly)
158 stationary can be accepted.

159 *2.4. Data analysis*

160 *2.4.1. Stride Interval Analysis*

161 Stride intervals were calculated using an automatic stride interval extrac-
162 tion algorithm (Chau and Rizvi, 2002). From the set of probabilistic stride
163 intervals, strides that fell outside the 0.01 and 99.99% of a gamma distribu-
164 tion fit were removed as these stride times were considered unphysiologically
165 long or short. It should be also mentioned that stride interval time series
166 were trimmed to 15 minutes duration. This was done by removing the first
167 59 seconds of data and only including data up to the 959th second. This
168 trimming was done to remove any extraneous static portions of the record-
169 ings and to exclude any effects due to acceleration towards the preferred
170 walking speed at the beginning of the trial and the deceleration at the end
171 of the walking trial.

172 *2.4.2. Stationarity Testing*

173 Due to the sensitivity of the RAT to window size, we tested stationarity
174 at a range of window sizes (10 to 45 strides), at increments of 5 strides. The
175 minimum window size was constrained to at least 10 stride intervals and the
176 maximum window size set such that a minimum of 10 windows were available
177 for analysis. This maintained an adequate number of data points as required
178 to estimate a single statistical parameter (Chau et al., 2005) when calculating
179 both the mean squared value within each interval and the total number of
180 reverse arrangements. As the time series lengths were not exact multiples
181 of the chosen window sizes, both ends of the time series were trimmed, as
182 justified earlier.

183 We investigated the sources of nonstationarity in all trials that violated

184 the hypothesis of stationarity. We divided each stride interval time series
185 into the chosen window length and the summary statistics of mean and vari-
186 ance were computed for each window. The null hypothesis of time invariance
187 of each summary statistic was tested via regression analysis. A 5% signifi-
188 cance level was used throughout. Left and right foot data were considered
189 separately.

190 **3. Results**

191 *3.1. Effect of window size*

192 The boxplot in Figure 1 shows the variation in the stationarity test statis-
193 tic, z_T , with different window sizes for the first music and TV sessions, for
194 the right foot. Similar plots were observed for all other walking trials, and
195 for the left foot. A few general observations are in order.

- 196 1. The stride interval time series were overwhelmingly nonstationary. The
197 median values of the stationarity test statistic generally fell outside
198 the stationary range at the 5% level of significance (denoted by the
199 dashed lines), i.e., usually $\text{med}|z_T| > 1.96$. Two trials exhibited weakly
200 stationary stride time series: (i) OW-NoMusic2 in music session 1, at
201 window sizes greater than 20 in the right foot and all window sizes for
202 the left foot (not shown), and (ii) OW-NoMusic1 in music session 2
203 at all window sizes for both feet (not shown). These two time series
204 constituted only 8.3% of the 24 unique combinations of session (music
205 or TV), trial (1,2 or 3) and side (left or right).
- 206 2. The stationarity test statistic alone did not suggest a preferred window
207 size. A Kruskal-Wallis test found that the mean ranks for the station-

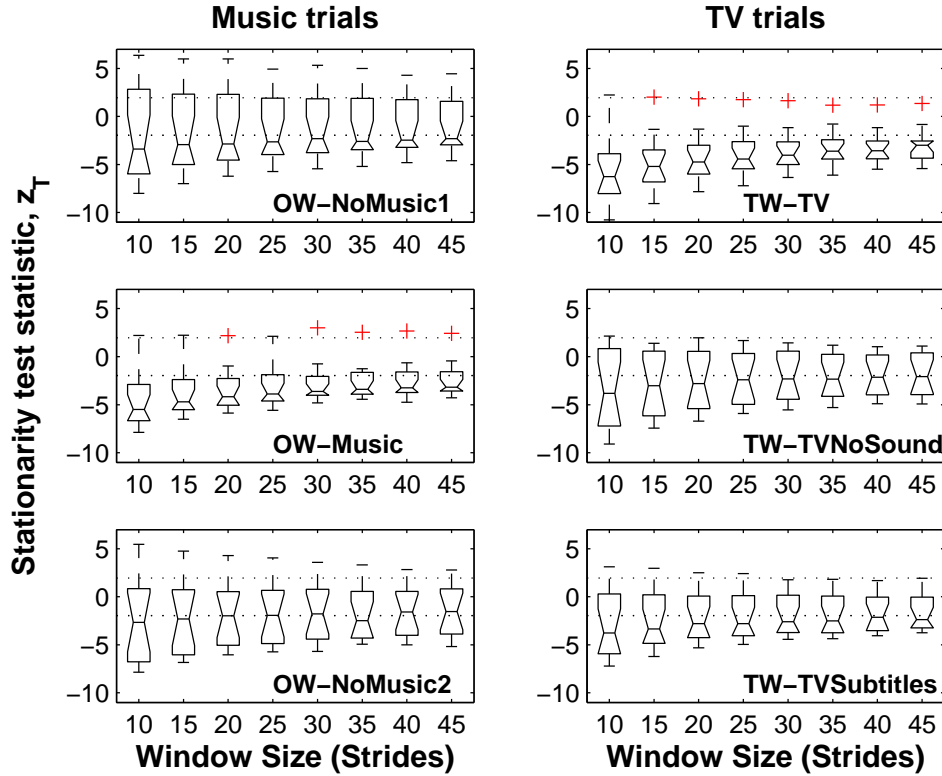


Figure 1: The effect of window size on the stationarity test statistic, z_T , for the first music (left panel) and TV (right panel) sessions, right foot.

208 arity test statistic were not significantly different among window sizes
 209 ($p \geq 0.95$) in all cases other than the two time series identified above.
 210 3. The number of nonstationary time series in each walking condition
 211 tended to decrease with increasing window size. At larger window sizes,
 212 a time series is divided into fewer segments; thus, a shorter mean square
 213 sequence is created, resulting in fewer comparisons between subsequent
 214 mean square values, and thereby reducing the number of opportunities
 215 for detecting a reverse arrangement.

216 4. The variation in the stationarity statistic decreased with increasing
217 window size. This trend is apparent in the progressive shortening of
218 the boxplot whiskers in Figure 1.

219 A reduction in the number of nonstationary stride series with larger win-
220 dow sizes is intuitively correct because even a slow varying trend would ap-
221 pear stationary at a sufficiently large window size. Hence the reverse arrange-
222 ments test is least reliable at larger window sizes where nonstationarities due
223 to fast varying trends may go undetected. This results in a trade-off between
224 maintaining an adequate number of stride intervals within each window and
225 generating a sufficiently long mean square sequence for identification of non-
226 stationarities; we chose an intermediate value of 25 stride intervals for the
227 window size for all further analysis. This resulted in 263 nonstationarities
228 in total: 124 in session 1 (31 per side for music and TV stimuli), and 139 in
229 session 2 (music: 34 (right), 33 (left); TV: 35 (right), 37 (left)).

230 3.2. *Between sessions and between sides*

231 Gait speeds were consistent among all trials ($p > 0.21$, Kruskal-Wallis
232 Test) and between sessions ($p > 0.12$, Mann-Whitney U-test) in the music
233 protocol. Similarly, in the TV protocol, no differences in gait speed among
234 trials ($p > 0.51$, Kruskal-Wallis Test) or between sessions ($p > 0.5$, Mann-
235 Whitney U-test) were observed. At the chosen window size of 25 and a given
236 side (left or right), there were no statistical differences in the stationarity test
237 statistic (z_T) values between sessions 1 and 2 for all corresponding music ($p >$
238 0.36 , Mann-Whitney U-test) trials. The same held true for all corresponding
239 TV ($p > 0.58$, Mann-Whitney U-test) trials. Further, z_T values did not differ

240 between left and right sides within any music ($p > 0.76$, Mann-Whitney U-
241 test) or TV trial ($p > 0.89$, Mann-Whitney U-test).

242 *3.3. Effect of music or TV on stride interval stationarity*

243 As seen in Figure 2, OW-Music (overground walking while listening to
244 music) generally had the highest number of nonstationarities out of the three
245 walks in the music sessions. A one-way ANOVA confirmed that the num-
246 ber of nonstationarities were different across conditions for both sessions
247 ($p < 10^{-4}$). Post-hoc pairwise comparisons revealed that in session 1, the
248 music and second no music walks differed significantly using a Bonferroni-
249 corrected adjusted significance level of 0.0167 ($p = 0.0017$, T-test) while
250 in session 2, the music and first no music walks were significantly different
251 ($p = 10^{-4}$). TW-TV (TV with sound) had the most nonstationarities of all
252 walking conditions. The trial (either TW-TVNoSound or TW-TVSubtitles)
253 with the least amount of nonstationarities varied with window size and ses-
254 sion. A one-way ANOVA verified significant differences across TV trials for
255 both sessions ($p < 6.2 \times 10^{-5}$) while post-hoc pairwise comparisons identified
256 significant differences between the TV with sound walk and the other two
257 TV walks ($p < 0.002$).

258 *3.4. Sources of non-stationarity*

259 As depicted by Figures 3 and 4, nonstationarities were mainly due to
260 variations in the mean stride interval over time. No stride series demonstrated
261 time-dependent variance alone, though 10-20% of nonstationarities could be
262 attributed to changes in both mean and variance in at least one walk in
263 all sessions. Also, in any one session, no more than one walk would have

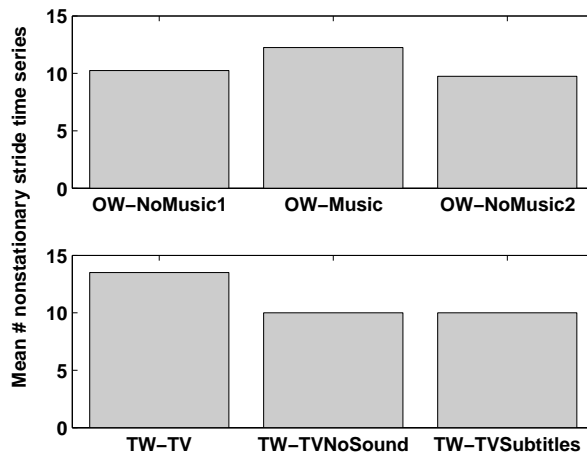


Figure 2: Average number of nonstationary time series per condition

264 nonstationarities due to unknown causes, and this was always less than 10%
 265 of the nonstationarities for that walk. Hence, for all walking conditions, the
 266 major cause of nonstationarity was a time-varying mean.

267 4. Discussion

268 4.1. Locomotor perspective of nonstationarity

269 The nonstationarities uncovered for each walking condition in this ex-
 270 periment are suggestive of the complexities inherent to gait. Even during
 271 overground walking at one’s self-selected comfortable walking speed, time-
 272 varying changes in the stride interval time series occur within a 15 minute
 273 walk and are largely due to a time-dependent mean. These changes in time
 274 may be the result of fatigue, loss or modification of concentration, anticipa-
 275 tion of task completion or boredom (Hausdorff et al., 1999). On the other
 276 hand, these nonstationarities may also be a result of intrinsic nonlinearities

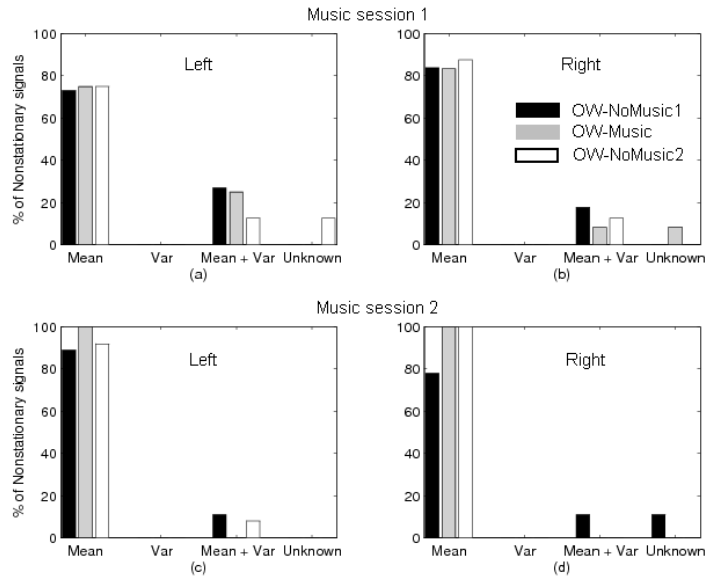


Figure 3: Sources contributing to nonstationarity of the time series during music sessions, as a percentage of the nonstationarities identified within each particular walking condition: left foot during first (a) and second (c) sessions; right foot during first (b) and second (d) sessions.

277 present in stride dynamics. Studies have shown that the healthy adult lo-
 278 comotor system possesses memory, that is, the change from one stride to
 279 the next contains a temporal structure that has been associated with statis-
 280 tical persistence (Hausdorff et al., 1999; Ghafari et al., 2009). Nonstation-
 281 arities may also arise, at least in part, as a consequence of the integration
 282 of multiscale information from various sensory inputs such as the visual,
 283 vestibular and proprioceptive systems (Dietz, 2002; West and Griffin, 1999).
 284 Furthermore, gait is achieved via the neuronal control system (Ghafari et al.,
 285 2009), requiring at least 30 major muscles working in concert, both tempo-
 286 rally and kinetically (Kwak, 2007). These interactions between the neuronal

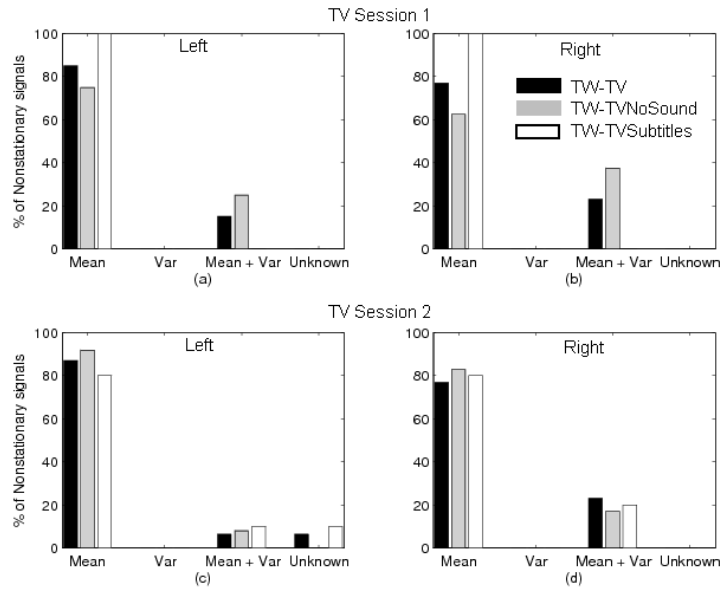


Figure 4: Sources contributing to nonstationarity of the time series, as a percentage of the nonstationarities identified within each particular walking condition: left foot during first (a) and second (c) sessions; right foot during first (b) and second (d) sessions.

287 and musculoskeletal systems can be affected by tactile, auditory and visual
 288 cues (Schauer and Mauritz, 2003; Prokop et al., 1997).

289 4.2. The effect of music

290 The results of the music sessions suggest that music does in fact alter
 291 one's stride interval dynamics. Overground walking with music led to more
 292 nonstationary stride interval time series than overground walking without
 293 music. During interviews after each session, nine of the fifteen participants
 294 commented that the music altered their strides. These participants reported
 295 that they walked to the tempo of the music or that the music caused them to
 296 lose concentration during the task. Indeed, previous studies have shown that
 297 music can increase walking speed and induce synchronization of walking pace

298 to the beat of the music (Styns et al., 2007). The music in these sessions
299 consisted of a random mix of popular songs, each with a different tempo.
300 Participants may have altered their stride rhythm to align with each tempo,
301 thus leading to more nonstationarities in the stride intervals.

302 Multiple studies have investigated the effect of rhythmic auditory stim-
303 ulation (RAS), that is, the effect of rhythmic stimuli, such as rock music
304 on human gait dynamics (Thaut et al., 1999). RAS is a neurologic tech-
305 nique that exploits the physiological effects of an auditory rhythm on the
306 motor system in order to improve the control of movement (Prassas et al.,
307 1997). In fact, RAS has been shown to improve gait performance in Parkin-
308 son’s disease (PD), Huntington’s chorea and hemiparesis (Schauer and Mau-
309 ritz, 2003; Hayashi et al., 2006). By entrainment through the reticulospinal
310 pathway, RAS can act as an internalized time keeper in rhythmic patterned
311 movement (Kwak, 2007; Prassas et al., 1997), intuitively, producing more
312 stationary gait. Contrary to this notion, nonstationarities during music lis-
313 tening increased in our study, presumably because gait is highly susceptible
314 to rhythmic sensory cuing (Thaut et al., 1999) and rhythmic changes in the
315 music occurred throughout the walk.

316 The magnification of nonstationary behaviour during the music walk
317 (OW-Music) may also be connected to a “distraction effect” that music can
318 cause during exercise at submaximal intensities (Yamashita et al., 2006).
319 Distraction may arise with increasing comfort level of the person exercising,
320 or a decrease in fatigue-induced stress (Yamashita et al., 2006) over time.
321 This distraction effect may also be attributed to a narrowing of one’s atten-
322 tion to focus only on the music and a reduction in the awareness of bodily

323 sensations such as fatigue (Karageorghis et al., 2009). Consequently, the
324 person perceives lower exertion during exercise (Karageorghis et al., 2009),
325 an effect linked to a reduction in the metabolic cost of exercise achieved by
326 promoting greater neuromuscular efficiency (Yamashita et al., 2006). The
327 fact that participants became cognizant of their distracted state implies that
328 they may have attempted to consciously correct their gait rhythm from time
329 to time, exacerbating the overall nonstationarity of the stride interval time
330 series. In such sense, the distraction might be more aptly described as cyclic
331 or intermittent rather than constant.

332 The combined effects of rhythmic synchronization and cyclic distraction,
333 combined with the distributed effect of music on the brain and motor control
334 system (Karageorghis et al., 2009; Emery et al., 2003) potentially height-
335 ened the nonstationarities observed in the walking with music trials through
336 increased central modulation of gait.

337 *4.3. The effect of TV and treadmill*

338 Intuitively, one might think that treadmill walking would result in more
339 stationary gait due to the constant speed constraint. However, our results
340 show the opposite to be true. Increased incidence of nonstationary stride in-
341 terval time series during treadmill walking may be the result of many factors.
342 The treadmill modality constrains walking by forcing the subject to maintain
343 a relatively constant speed as well as restricting the area of movement and
344 the subject's stride length, all factors shown to influence human locomotor
345 control (Wall and Charteris, 1981). When subjects first start the trial on the
346 treadmill, they must adjust their speed and stride length to that permitted
347 by the moving belt. In particular, it has been found that stride intervals

348 increase over time while walking on a treadmill (Wall and Charteris, 1981),
349 which may in part explain the time-varying mean. During the television
350 trials, participants visually attended to the TV program while attempting
351 to walk at a steady pace. Studies have shown that the stride length is in-
352 fluenced by visual distractions such as changes in optic flow (Prokop et al.,
353 1997). Most of the participants (87%) subjectively acknowledged that watch-
354 ing television altered their gait. Many stated that the TV diminished their
355 focus on walking, resulting in near stumbles on multiple occasions. Others
356 thought the television helped them to concentrate on their walking, so as to
357 mitigate the risk of falling.

358 One of the television trials involved listening to the program. Like the
359 walking with music trial, listening to the television further diverted the par-
360 ticipant’s attention from the walking task at hand, possibly causing nonsta-
361 tionarities through the distraction effect introduced above. Based on evidence
362 from a mobile phone study (Murray-Smith et al., 2007), it is also possible that
363 people walked in step to conversations on the television. Whether television
364 caused increased concentration, distraction or synchronization, the combina-
365 tion of visual and auditory distractions seems to have led to additional central
366 modulations of the locomotor system, resulting in more nonstationarities in
367 the stride interval time series.

368 *4.4. Relevance to pedestrian mobility*

369 Interestingly, in a recent study of pedestrian safety while listening to
370 music, Renfroe et al. (2010) noted that self-reported distraction levels were
371 highest when participants listened to music that changed from a slow to fast
372 tempo. Indeed, listening to headphones while walking has been implicated

373 as a distraction that leads to fewer cautionary behaviors (Bungum et al.,
374 2005). As this study deployed mixed musical stimuli with slow to fast tempo
375 transitions, exaggerated nonstationarities may be an indicator of greater dis-
376 traction, and ultimately reduced safety. Using television as an audio and
377 visual distraction, Malone and Bastian (2010) found that distracted partici-
378 pants exhibited slower adaptation of gait in response to perturbations by a
379 split belt treadmill. Our results offer a potential explanation. With the dis-
380 traction of television, more stride interval time series became nonstationary,
381 suggesting the occurrence of more locomotor adaptations. It is thus plausible
382 that as the locomotor system becomes preoccupied with adaptations asso-
383 ciated with the television distraction, its ability to react to perturbations
384 diminishes.

385 *4.5. Relevance to analysis of statistical persistence*

386 This investigation has determined that the stride interval time series of
387 adults walking to music or watching television are generally nonstationary.
388 Thus when performing further analyses on stride time series, such as the
389 estimation of the fractal scaling exponent, which is beyond the scope of the
390 current manuscript, it is important to use techniques that account for this
391 nonstationary behaviour. In particular, wavelet-based approaches (e.g., Si-
392 monsen et al. (1998)) or DFA in some cases (e.g., Peng et al. (1994); Chen
393 et al. (2002)) are particularly suited to the detection of statistical persistence
394 in time series which exhibit nonstationary behaviour (Peng et al., 1994; Chen
395 et al., 2002). However, one should be careful when applying DFA to non-
396 stationary time series since certain nonstationarities such as random spikes
397 or time-evolving correlations can differentially influence the scaling estimate

398 (e.g., introduce cross-overs at certain scales) (Chen et al., 2002).

399 **5. Conclusion**

400 In this paper, we investigated the stationarity of stride intervals in healthy
401 adults while exposed to several conditions of music and television stimuli.
402 Overall, the stride interval time series were nonstationary under these condi-
403 tions. Nonstationarities, primarily in the form of time-varying means, were
404 particularly abundant when walking on a treadmill while watching television
405 with sound. The heightened number of nonstationarities in trials involving
406 auditory stimuli suggests that natural gait rhythms are especially suscepti-
407 ble to the effects of music and television sounds. The consequence of altered
408 stride dynamics on environmental vigilance (e.g., pedestrian safety) and re-
409 sponsiveness to perturbations (e.g., risk of falls) ought to be investigated in
410 future research.

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